

## Exam in Statistical Methods, 2017-10-25

**Time allowed:** kl: 8-12

**Allowed aids:** Calculator. One handwritten A4 paper (both sides) with the students own notes.

**Assisting teacher:** Hector Rodriguez-Deniz

**Grades:** A=19-20 points, B=17-18p, C=14-16p, D=12-13p, E=10-11p

Provide a detailed report that shows motivation of the results.

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### 1

Let a uniformly distributed random variable  $Y$  with parameters  $\theta_1, \theta_2$  such that

$$f(y) = \begin{cases} \frac{1}{\theta_2 - \theta_1}, & \theta_1 \leq y \leq \theta_2 \\ 0, & \text{elsewhere} \end{cases} \quad \text{be the density function of } Y.$$

- Show that  $E[Y] = \frac{\theta_1 + \theta_2}{2}$ . 1p
- Find the cumulative distribution function  $F(y)$ , and show that  $F(y) = y$  when  $Y \sim \text{Uniform}(0,1)$ . 1p
- Calculate the probability  $P(5 \leq Y \leq 10)$  when  $Y \sim \text{Uniform}(0,20)$ . 1p

### 2

Let the bivariate random variable  $(Y_1, Y_2)$  have joint density function:

$$f(y_1, y_2) = \begin{cases} cy_1y_2, & 0 \leq y_2 \leq y_1 \leq 1 \\ 0, & \text{elsewhere} \end{cases}$$

- Find  $c$  so that  $f(y_1, y_2)$  is a density function. 2p
- Find the marginal distribution  $f(y_1)$  1p
- Find  $P(Y_1 > 0.5, Y_2 > 0.5)$ . 2p

### 3

Let  $Y_1, Y_2, \dots, Y_n$  be a random sample from a normal distribution with mean  $\mu$  and variance 1. The

density is given by  $f(y|\mu) = \frac{1}{\sqrt{2\pi}} e^{-0.5(y-\mu)^2}$

- Estimate  $\mu$  using Maximum Likelihood. 3p
- Show that  $\hat{\mu}_{MLE}$  is an unbiased and consistent estimator of  $\mu$ . 2p

## 4

We collect the following random sample  $(y_i, x_i)$  of size  $n = 7$  where  $Y$  is the dependent (output) variable and  $X$  the independent (input) variable.

| Y | X  |
|---|----|
| 5 | -2 |
| 4 | -1 |
| 4 | 0  |
| 3 | 1  |
| 2 | 2  |
| 1 | 3  |
| 1 | 4  |

- Set up the simple linear regression model and estimate  $\beta_0$  and  $\beta_1$ . 2p
- Test if the slope is zero. You may use without showing any calculations that  $SSE = 0,571$ . Use 5% significance level. Interpret your result. 1p
- Find a 90% confidence interval for  $E[Y] = \beta_0 + \beta_1 x^*$  at  $x^* = 5$ . 1p

## 5

We want to estimate the number of patients arriving in an emergency room at Linköping's Hospital during 22:00-23:00 on Fridays. We collect a sample of size  $n$ , and we use a Poisson distribution for modeling the data, i.e.  $Y \sim \text{Poisson}(\lambda)$ . Now we place a Gamma prior on  $\lambda$ , i.e.  $\lambda \sim \text{Gamma}(\alpha, \beta)$ .

- Find the posterior distribution of  $\lambda$ , including the proportionality constant. 2p
- If we collect a sample such that  $n = 10, \sum y = 71$ , what would be your Bayes estimate for the number of patients after using a prior  $\text{Gamma}(4, 2)$ ? 1p

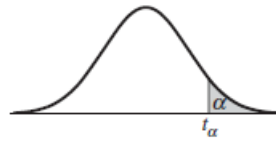
### Discrete Distributions

| Distribution      | Probability Function   | Mean           | Variance   |
|-------------------|--|----------------|--|
| Binomial          | $p(y) = \binom{n}{y} p^y (1-p)^{n-y};$ $y = 0, 1, \dots, n$  | $np$           | $np(1-p)$  |
| Geometric         | $p(y) = p(1-p)^{y-1};$ $y = 1, 2, \dots$   | $\frac{1}{p}$  | $\frac{1-p}{p^2}$  |
| Hypergeometric    | $p(y) = \frac{\binom{r}{y} \binom{N-r}{n-y}}{\binom{N}{n}};$ $y = 0, 1, \dots, n \text{ if } n \leq r,$ $y = 0, 1, \dots, r \text{ if } n > r$ | $\frac{nr}{N}$ | $n \left( \frac{r}{N} \right) \left( \frac{N-r}{N} \right) \left( \frac{N-n}{N-1} \right)$ |
| Poisson           | $p(y) = \frac{\lambda^y e^{-\lambda}}{y!};$ $y = 0, 1, 2, \dots$   | $\lambda$      | $\lambda$  |
| Negative binomial | $p(y) = \binom{y-1}{r-1} p^r (1-p)^{y-r};$ $y = r, r+1, \dots$   | $\frac{r}{p}$  | $\frac{r(1-p)}{p^2}$   |

### Continuous Distributions

| Distribution | Probability Function  | Mean                            | Variance   |
|--------------|---|---------------------------------|--|
| Uniform      | $f(y) = \frac{1}{\theta_2 - \theta_1}; \theta_1 \leq y \leq \theta_2$   | $\frac{\theta_1 + \theta_2}{2}$ | $\frac{(\theta_2 - \theta_1)^2}{12}$                         |
| Normal       | $f(y) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left[-\left(\frac{1}{2\sigma^2}\right)(y - \mu)^2\right]$<br>$-\infty < y < +\infty$   | $\mu$                           | $\sigma^2$   |
| Exponential  | $f(y) = \frac{1}{\beta} e^{-y/\beta}; \beta > 0$<br>$0 < y < \infty$  | $\beta$                         | $\beta^2$  |
| Gamma        | $f(y) = \left[\frac{1}{\Gamma(\alpha)\beta^\alpha}\right] y^{\alpha-1} e^{-y/\beta};$<br>$0 < y < \infty$                       | $\alpha\beta$                   | $\alpha\beta^2$  |
| Chi-square   | $f(y) = \frac{(y)^{(v/2)-1} e^{-y/2}}{2^{v/2} \Gamma(v/2)}$<br>$y^2 > 0$  | $v$                             | $2v$   |
| Beta         | $f(y) = \left[\frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)}\right] y^{\alpha-1} (1 - y)^{\beta-1};$<br>$0 < y < 1$ | $\frac{\alpha}{\alpha + \beta}$ | $\frac{\alpha\beta}{(\alpha + \beta)^2(\alpha + \beta + 1)}$ |

Table 5 Percentage Points of the  $t$  Distributions



| $t_{.100}$ | $t_{.050}$ | $t_{.025}$ | $t_{.010}$ | $t_{.005}$ | df   |
|------------|------------|------------|------------|------------|------|
| 3.078      | 6.314      | 12.706     | 31.821     | 63.657     | 1    |
| 1.886      | 2.920      | 4.303      | 6.965      | 9.925      | 2    |
| 1.638      | 2.353      | 3.182      | 4.541      | 5.841      | 3    |
| 1.533      | 2.132      | 2.776      | 3.747      | 4.604      | 4    |
| 1.476      | 2.015      | 2.571      | 3.365      | 4.032      | 5    |
| 1.440      | 1.943      | 2.447      | 3.143      | 3.707      | 6    |
| 1.415      | 1.895      | 2.365      | 2.998      | 3.499      | 7    |
| 1.397      | 1.860      | 2.306      | 2.896      | 3.355      | 8    |
| 1.383      | 1.833      | 2.262      | 2.821      | 3.250      | 9    |
| 1.372      | 1.812      | 2.228      | 2.764      | 3.169      | 10   |
| 1.363      | 1.796      | 2.201      | 2.718      | 3.106      | 11   |
| 1.356      | 1.782      | 2.179      | 2.681      | 3.055      | 12   |
| 1.350      | 1.771      | 2.160      | 2.650      | 3.012      | 13   |
| 1.345      | 1.761      | 2.145      | 2.624      | 2.977      | 14   |
| 1.341      | 1.753      | 2.131      | 2.602      | 2.947      | 15   |
| 1.337      | 1.746      | 2.120      | 2.583      | 2.921      | 16   |
| 1.333      | 1.740      | 2.110      | 2.567      | 2.898      | 17   |
| 1.330      | 1.734      | 2.101      | 2.552      | 2.878      | 18   |
| 1.328      | 1.729      | 2.093      | 2.539      | 2.861      | 19   |
| 1.325      | 1.725      | 2.086      | 2.528      | 2.845      | 20   |
| 1.323      | 1.721      | 2.080      | 2.518      | 2.831      | 21   |
| 1.321      | 1.717      | 2.074      | 2.508      | 2.819      | 22   |
| 1.319      | 1.714      | 2.069      | 2.500      | 2.807      | 23   |
| 1.318      | 1.711      | 2.064      | 2.492      | 2.797      | 24   |
| 1.316      | 1.708      | 2.060      | 2.485      | 2.787      | 25   |
| 1.315      | 1.706      | 2.056      | 2.479      | 2.779      | 26   |
| 1.314      | 1.703      | 2.052      | 2.473      | 2.771      | 27   |
| 1.313      | 1.701      | 2.048      | 2.467      | 2.763      | 28   |
| 1.311      | 1.699      | 2.045      | 2.462      | 2.756      | 29   |
| 1.282      | 1.645      | 1.960      | 2.326      | 2.576      | inf. |